

Lecture 22:

Pseudorandomness &

Unpredictability

Another notion of randomness?

Unpredictability

Def. $X = X_1 \dots X_n$ is "next bit unpredictable"
 if \forall ppt P , \exists negligible fctn $\epsilon(n)$
 s.t. $\Pr_{X, i \in_R [n], \text{coins of } P} [P(X_1, \dots, X_{i-1}) = X_i] \leq \frac{1}{2} + \epsilon(n)$

Note:

- x uniform $\Rightarrow \epsilon = 0$
 - x statistically close to uniform $\Rightarrow \epsilon(n)$ negligible
 i.e. ϵ -close for ϵ negligible
 - x indistinguishable from uniform $\Rightarrow \epsilon(n)$ negligible
 (by ppt)
- since: "predict next bit" is a statistical test
 if can pass \Rightarrow can distinguish from uniform
 what about other direction?

Cool theorem: next bit unpredictability + pseudorandomness
 are equivalent!!

Thm X is p.r. iff X is n.b.u.

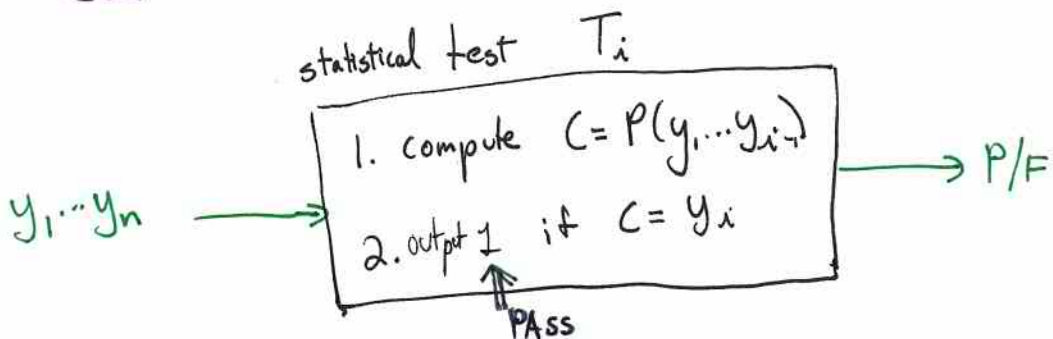
Pf.

\Rightarrow (show X not nbu $\Rightarrow X$ not p.r.)

Assume $\exists k$ st. $\Pr_{X, i \in [n], \text{coins}} [P(X_1, \dots, X_{i-1}) = X_i] \geq \frac{1}{2} + \frac{1}{n^k}$

then $\exists i$ st. $\Pr_{X, \text{coins}} [P(X_1, \dots, X_{i-1}) = X_i] \geq \frac{1}{2} + \frac{1}{n^k}$

Construct stat test T_i distinguishing X from U



$$\Pr_{y_1 \dots y_n \in X} [T_i(y_1 \dots y_n) = 1] \geq \frac{1}{2} + \frac{1}{n^k}$$

$$\Pr_{y_1 \dots y_n \in U_n} [T_i(y_1 \dots y_n) = 1] = \frac{1}{2}$$

} more likely to "pass" n.b. test
 \Downarrow
 statistical test that distinguishes

$\therefore X$ is not p.r.

← (X not p.r. $\Rightarrow \exists$ n.b. test
for some i)

if X not p.r.

$$\exists \text{ ppt } T \text{ st. } \left| \Pr_{x \in X} [T(x)=1] - \Pr_{u \in U} [T(u)=1] \right| > \frac{1}{n^k}$$

wlog advantage > 0
else, use \bar{T}

for infinitely many n

use hybrid argument to construct n.b. predictor:

$$D_0 \equiv U \equiv u_1 \dots u_n$$

$$D_1 = X_1 u_2 \dots u_n$$

$$D_2 = X_1 X_2 u_3 \dots u_n$$

\vdots

$$D_n = X_1 \dots X_n \equiv X$$

Main idea:

run statistical test

if says "pseudorandom" output the bit

else, seems "not right" so output the complement of the bit

← i.e. outputs "1"

some preliminary calculations:

$$\begin{aligned} \frac{1}{n^k} &< \Pr_{x \in D_n} [T(x)=1] - \Pr_{x \in D_0} [T(x)=1] \\ &= \sum_{i=1}^n \Pr_{x \in D_i} [T(x)=1] - \Pr_{x \in D_{i-1}} [T(x)=1] \end{aligned} \quad \text{telescoping}$$

divide by n :

$$\frac{1}{n^{k+1}} < \frac{1}{n} \sum_{i=1}^n \Pr_{x \in D_i} [T(x)=1] - \Pr_{x \in D_{i-1}} [T(x)=1]$$

so: $\exists i$ s.t. $\Pr_{x \in D_i} [T(x)=1] - \Pr_{x \in D_{i-1}} [T(x)=1] \geq \frac{1}{n^{k+1}}$

so define next bit predictor for this i :

n.b. p_i for i :

$P(x_1 \dots x_{i-1})$:

1. choose $u_1 \dots u_n \in_K \{0,1\}^{n-i}$

2. $b \leftarrow T(x_1 \dots x_{i-1} u_1 \dots u_n)$

3. if $b=1$ output u_i
else output \bar{u}_i

$\leftarrow u_i$ seems right
ie. $x_1 \dots x_{i-1} u_1 \dots u_n$
looks like output
of PRG

Note:

$P(x_1 \dots x_{i-1}) = x_i$

iff

\downarrow T passes u_i

$b=1 \quad \wedge \quad u_i = x_i$

or $b=0 \quad \wedge \quad \bar{u}_i = x_i$

\uparrow

T fails u_i

$$\begin{aligned}
 \Pr [\underbrace{P(x_1, \dots, x_{i-1}) = x_i}_{\text{event "x"}}] &= \Pr [* | u_i = x_i] \cdot \underbrace{\Pr [u_i = x_i]}_{= 1/2} + \Pr [* | u_i \neq x_i] \cdot \underbrace{\Pr [u_i \neq x_i]}_{= 1/2} \\
 &= \frac{1}{2} \cdot [\Pr [b=1 | u_i = x_i] + \Pr [b=0 | u_i \neq x_i]] \\
 &\quad \uparrow \qquad \qquad \qquad \uparrow \\
 &\quad T \text{ "passes"} \qquad \qquad T \text{ "fails"} \\
 &\qquad \qquad \qquad \qquad \qquad \qquad = 1 - \Pr [b=1 | u_i \neq x_i] \\
 &= \frac{1}{2} [1 + \Pr [T(x_1, \dots, x_{i-1}, x_i, u_{i+1}, \dots, u_n) = 1] - \Pr [T(x_1, \dots, x_{i-1}, \bar{x}_i, u_{i+1}, \dots, u_n) = 1]]
 \end{aligned}$$

add +
subtract (b) →

$$\begin{aligned}
 &= \frac{1}{2} + \frac{1}{2} [\underbrace{\Pr [T(x_1, \dots, x_{i-1}, x_i, u_{i+1}, \dots, u_n) = 1]}_{D_i} - \underbrace{\Pr [T(x_1, \dots, x_{i-1}, \bar{x}_i, u_{i+1}, \dots, u_n) = 1]}_{D_{i-1}}] \\
 &+ \frac{1}{2} [\underbrace{\Pr [T(x_1, \dots, x_{i-1}, u_i, u_{i+1}, \dots, u_n) = 1]}_{(b)} - \underbrace{\Pr [T(x_1, \dots, x_{i-1}, \bar{x}_i, u_{i+1}, \dots, u_n) = 1]}_{(c)}]
 \end{aligned}$$

what is this?

see ** calculation - it = (a) - (b)!

$$\cong \frac{1}{2} + \frac{1}{2} \cdot 2 \cdot \frac{1}{n^{k+1}} = \frac{1}{2} + \frac{1}{n^{k+1}}$$

$$** : \quad \Pr [T(x_1, \dots, x_{i-1}, u_i, \dots, u_n) = 1] = \frac{\Pr [T(x_1, \dots, x_{i-1}, u_{i+1}, \dots, u_n) = 1] + \Pr [T(x_1, \dots, \bar{x}_i, u_{i+1}, \dots, u_n) = 1]}{2}$$

$$(b) = \frac{(a) + (c)}{2} \Rightarrow (a) - (b) = (b) - (c)$$

Curious Note:

Since order of bits irrelevant for PRG statistical test T,
order of bits irrelevant for prediction.

so $b_1 \dots b_n$ hard for s.b. test $\Rightarrow b_n \dots b_1$ hard for n.b. test.

How do we construct PRG's

want to use "computational hardness"
will start with 1-way functions.

def. f is 1-way if

1) f computable in deterministic ptime

2) \forall ppt A , \exists negligible $\epsilon(n)$

st. $\forall n$ big enough

$$\Pr_{x, \text{coins of } A} [A(f(x)) \in \underbrace{f^{-1}(f(x))}_{\text{any inverse of } f(x)}] \leq \epsilon(n)$$

notes:

1) A ptime in n , not just $|f(x)|$ (which might be small)

2) don't need to find "right" inverse -
just some inverse

Why is this good? We have candidates!

1) $f(x,y) = x \cdot y$ factoring

2) $f_{m,e}(x) = x^e \pmod m$

RSA $m=pq$

3) $f_m(x) = x^2 \pmod m$

Rabin's fctn

(square roots mod m)

4) $f_{p,q}(x) = g^x \pmod p$

discrete log

For PRGs to exist, 1-way fctns must exist:

Claim $f: \{0,1\}^n \rightarrow \{0,1\}^{2n}$ PRG $\Rightarrow f$ is o.w.

idea if f not o.w.

\exists inverter algorithm which sometimes works
use it to give statistical test on PRG output
(i.e. output "1" if finds inverse)

- sometimes works on PRG output
- almost never works on truly random bits

But do 1-way fctns imply existence of PRG's?

Yes, but much harder to prove...

Thm [HILL] 1-way fctn exist \Rightarrow PRG's exist

so PRG's exist
 \Downarrow
 1-w.f.'s exist!

here, a weaker result

Thm 1-way permutations exist \Rightarrow PRG's exist
 fctn that is 1-1 + onto
 so preimages are unique