

Lecture 20:

Pseudorandomness

Pseudorandom Generators

Given n random bits, generate $m \gg n$ bits that "look random"



What is "random looking"?

- L_1 distance to uniform is small ← impossible

- Useful for randomized algorithms

e.g. pairwise independent \leftarrow doesn't suffice for all algorithms
k-wise

- "Computational Indistinguishability" \leftarrow good for all ptime algs but based on complexity assumption
by ptime algorithms

- Kolmogorov complexity \approx incompressible \leftarrow Success is undecidable

Computational Indistinguishability

Sequence of bits $\rightarrow X$ "looks random" if no efficient (ptime, small space, low depth ckt...) can distinguish X from uniform Y , i.e.

statistical test \rightarrow algorithm

$$\Delta(X, Y) = \max_{T \text{ that is "efficiently computable"}} | \Pr[X \in T] - \Pr[Y \in T] |$$

is small
 need to define this

def. Computational Indistinguishability (C.I.)

Let X_n, Y_n be sequences of r.v.'s on $\{0,1\}^n \leftarrow$ or $\text{poly}(n)$

We say $\{X_n\}, \{Y_n\}$ are $\epsilon(n)$ -indistinguishable

for time $t(n)$ if \forall probabilistic poly time algorithm T running in time $t(n)$
 \downarrow test \downarrow or circuits? (nonuniform)

$$| \Pr[T(X_n)=1] - \Pr[T(Y_n)=1] | \leq \epsilon(n)$$

advantage of T

$\forall n$ large enough,

(Probabilities are over X_n, Y_n , coin tosses of T)

Comments

• if $\epsilon(n)$ not specified then $\epsilon(n) = \frac{1}{t(n)}$

• $X_n \stackrel{c}{=} Y_n$ is notation for C.I. (w/o ϵ)

if $\frac{1}{n^c}$ -indistin for time $n^c \forall c$

equivalently:

\forall ppt T , $\exists \epsilon(n) = n^{-\omega(1)}$ s.t.

$$| \Pr [T(X_n) = 1] - \Pr [T(Y_n) = 1] | \leq \epsilon(n)$$

• X_n, Y_n C.I. in nonuniform model time $t(n)$ (NCI)

if also holds when given $\leq t(n)$ advice bits
i.e. encode in circuit

Def. $\epsilon(n)$ "negligible" if $\epsilon(n) < \frac{1}{n^c} \forall c$

Def. "Pseudorandom" X_n is p.r. if $X_n \stackrel{c}{=} U_n$

A nice theorem: CI in nonuniform model \Rightarrow K reps CI in nonuniform model

Thm X_n, Y_n NCI
then $\forall K = \text{poly}(n)$ X_n^K, Y_n^K are NCI
 $\swarrow \searrow$
K independent copies

Pf. By induction on K

$H_i = X_n^{k-i} Y_n^i$ so $H_0 = X_n^k$
 $H_k = Y_n^k$ "hybrid distributions"

Consider $H_0 \dots H_k$:

$$H_0 = X_n X_n \dots X_n X_n$$

$$H_1 = X_n X_n \dots X_n Y_n$$

$$X_n X_n \dots X_n Y_n Y_n$$

⋮

$$X_n Y_n \dots Y_n$$

$$H_k = Y_n \dots Y_n$$

Assume for contradiction that \exists ppt T s.t.

$$|\Pr [T(X_n^K) = 1] - \Pr [T(Y_n^K) = 1]| > \epsilon$$

Then we'll construct ppt T' distinguishing $X_n \neq Y_n$
for infinitely many n 's
(if finite, there is largest n_0 for which doesn't work so no contradiction)

How?

Note $|\Pr [T(X_n^k)=1] - \Pr [T(Y_n^k)=1]|$
 $= \left| \sum_{i=1}^k (\Pr [T(H_{i-1})=1] - \Pr [T(H_i)=1]) \right|$ telescoping sum
 $> \epsilon$

$\Rightarrow \exists i$ st. $|\Pr [T(H_{i-1})=1] - \Pr [T(H_i)=1]| > \epsilon/k$
 $\Pr [T(X_n^{k-i} X_n Y_n^{i-1})=1] - \Pr [T(X_n^{k-i} Y_n Y_n^{i-1})=1]$

wlog assume absolute value positive

\Rightarrow via averaging,

$\exists X_1 \dots X_{k-i}, Y_{k-i+2} \dots Y_k$ st.

$\Pr [T(X_1 \dots X_{k-i} X_n Y_{k-i+2} \dots Y_k)=1]$
 $- \Pr [T(X_1 \dots X_{k-i} Y_n Y_{k-i+2} \dots Y_k)=1] > \epsilon/k$

So define $T'(z) = T(X_1 \dots X_{k-i} z Y_{k-i+2} \dots Y_k)$

So T' distinguishes X_n & Y_n with non negligible probability

- nonuniform - put $X_1 \dots X_{k-i}, Y_{k-i+2} \dots Y_k$ into circuit
- ie. many such ckts, one for each possible advice, & we are claiming at least one works so \exists poly size ckt which distinguishes
- time $(T') = \text{time}(T) + O(nk)$ \leftarrow extra time for advice
- $\epsilon' = \text{adv}(T') > \epsilon/k$

An almost nice theorem: CI in uniform model + sampleable \Rightarrow k reps CI in uniform model

prg 6
Sp2014

Thm.

if $X_n \equiv Y_n$ VCI (CI in uniform model)

+ X_n, Y_n ptime sampleable

(\exists ppt M s.t. $M(1^n) = X_n$)

then $X_n^k \equiv Y_n^k$

Pf. Assume \exists ptime T s.t. $\Pr[T(X_n)=1] - \Pr[T(Y_n)=1] > \epsilon$

$T'(z)$: choose $i \in_R \{1..k\}$ H_i
 output $T(X_n^{k-i} z Y_n^{i-1})$
 generate via M

if $X_n^k \not\equiv Y_n^k$ then

$$\Pr[T'(X_n)=1] - \Pr[T'(Y_n)=1]$$

$$= \left(\frac{1}{k} \sum_{i=1}^k \Pr[T(H_{i-1})=1] \right) - \left(\frac{1}{k} \sum_{i=1}^k \Pr[T(H_i)=1] \right)$$

$$= \frac{1}{k} \left[\Pr[T(H_0)=1] - \Pr[T(H_k)=1] \right]$$

$$> \frac{\epsilon}{k}$$

$$+ \text{time}(T') = \text{time}(T) + k(n) g(n)$$

\uparrow
time to sample X_n, Y_n

$\rightarrow \leftarrow$



Def. [Blum-Micali-Yao]

$G: \{0,1\}^{\ell(n)} \rightarrow \{0,1\}^n$ is a pseudorandom generator (PRG)

if (1) $\ell(n) < n$
(2) $G(U_{\ell(n)}) \stackrel{c}{=} U_n$

G is "efficient" if $\underbrace{\text{poly time}}_{\text{in } n}$ Computable since $\ell(n)$ could be as small as 1

Comments:

- must fool all prime statistical tests T , even those with runtime $\gg G$'s } useful for crypto
- can generalize to ϵ -PRG against nonuniform machines running in time $t(n)$
- create 1-time-pads
- recall earlier lecture:

method of enumeration (try all random seeds & take majority vote)

$$\Rightarrow \text{BPP} \subseteq \bigcup_c \text{PTIME}(2^{\ell(n)^c} n^c)$$

ie. if $\ell(n) = O(\log n)$ BPP = P

Can we prove existence of PRG's?

well, if we could, then we would also be able to prove $P \neq NP$!

Thm efficient PRG's exist $\Rightarrow P \neq NP$

Pf. (show $P = NP \Rightarrow$ no PRG's)
Assume $G: \{0,1\}^{\ell(n)} \rightarrow \{0,1\}^n$ + can be computed in ptime + $\ell(n) < n$
Given G let $(G \text{ is proposed efficient prg})$

$$T(x) = \begin{cases} 1 & \text{if } \exists y \text{ st. } G(y) = x \\ 0 & \text{o.w.} \end{cases}$$

note, $T \in NP$ since can guess y
+ verify via G
since G is ptime

$$\Pr_{x \in U_{\ell(n)}} [T(G(x)) = 1] = 1$$

$$\Pr_{y \in U_n} [T(y) = 1] \leq \frac{2^{\ell(n)}}{2^n} \leq \frac{1}{2} \quad \text{since } \ell(n) < n$$

if $P = NP$ then $T \in P$ + can distinguish the two distributions
but then G is not PRG $\rightarrow \leftarrow$

so need complexity theoretic assumptions to construct PRG's. ▣